Brownian Motion and Levy's Theorem

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Theorem (Levy's Theorem)

Let $X = \{X_t, \mathcal{F}_t, 0 \le t < \infty\}$ be a continuous, adapted process taking values in \mathbb{R} such that the process

$$M_t = X_t - X_0, \quad 0 \le t < \infty$$

is a **continuous local martingale** relative to the filtration $\{\mathcal{F}_t\}$, and whose **quadratic variation** is given by

 $\langle M, M \rangle_t = t;$

Then $\{X_t, \mathcal{F}_t, 0 \leq t < \infty\}$ is a Brownian Motion.





Levy's Continuity Theorem

► Convergence in Distribution of Random Variables ↔ Pointwise Convergence of their Characteristic Functions



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Characterization of the Brownian Motion



Levy's Continuity Theorem

- ► Convergence in Distribution of Random Variables ↔ Pointwise Convergence of their Characteristic Functions
- Characterization of the Brownian Motion
- Many Other Topics (For Example Infinitely Divisible Laws)

- Probability Space $(\Omega, \mathcal{F}, \mathbb{P})$
- Filtration {F_t, t ∈ Γ}, where Γ = ℕ (discrete time) or Γ = ℝ₀⁺ (continuous time)

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Definition (Usual Conditions)

A filtration \mathcal{F}_t is said to have the usual conditions if it is right continuous and \mathcal{F}_0 contains all the \mathbb{P} -negligible events in \mathcal{F} .

Definition (Measurable Process)

A stochastic process X is called measurable if, $\forall A \in \mathcal{B}(\mathbb{R}^d)$, the set $\{(t, \omega), X_t(\omega) \in A\}$ belongs to the product σ -field $\mathcal{B}(\mathbb{R}^+_0) \otimes \mathcal{F}$, or

$$(t,\omega) \to X_t(\omega) : (\mathbb{R}^+_0 \times \Omega, \mathcal{B}(\mathbb{R}^+_0) \otimes \mathcal{F}) \to (\mathbb{R}^d, \mathcal{B}(\mathbb{R}^d))$$

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is measurable.

Definition (Adapted Process)

A sequence of r.v. X_t is said to be adapted to \mathcal{F}_t if $X_t \in \mathcal{F}_t, \forall t \in \Gamma$, which means that for every $A \in \mathcal{B}(\mathbb{R}^d)$, $\{\omega \in \Omega : X_t(\omega) \in A\} \in \mathcal{F}_t$.

Conditional Expectation

Given a σ -field \mathcal{G} and an integrable random variable Y, $\mathbb{E}_{\mathcal{G}}(Y)$ must satisfy the conditions:

1. $\mathbb{E}_{\mathcal{G}}(Y) \in \mathcal{G}$ 2. $\forall \Lambda \in \mathcal{G}$

$$\int_{\Lambda} Y(w) \mathbb{P}(dw) = \int_{\Lambda} \mathbb{E}_{\mathcal{G}}(Y)(w) \mathbb{P}(dw).$$

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 (Geometric Perspective) Suppose E(Y²) < ∞. Then E_G(Y) is the random variable X ∈ G that minimizes E(Y − X)².

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To understand what information do we need from the past, our focus is to analyze $\mathbb{E}(S_{n+1}|S_n)$.

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- This independence condition is much stronger than what we need!
- Mathematically, martingales appeared to give exactly what we need! We only need that E(X_{n+1}|S_n) = 0.

Definition (Martingale)

A sequence of r.v. X_t is said to be a martingale with respect to a filtration $\{\mathcal{F}_t, t \in \mathbb{R}^+_0\}$ if it satisfies:

- 1. $\mathbb{E}(|X_t|) < \infty$ (integrability),
- 2. $X_t \in \mathcal{F}_t$ (measurability),

3.
$$\mathbb{E}_{\mathcal{F}_s}(X_t) = X_s, \quad \forall s \leq t.$$

Also, if in the last equation we have \leq or \geq , X_t is called a supermartingale or submartingale, respectively. The class of continuous martingales is denoted by \mathcal{M}^c .

Doob- Meyer Decomposition

Theorem (Doob-Meyer Decomposition)

Any submartingale X_t can be written in $X_t = M_t + A_t$, where $M_t \in \mathcal{M}^c$ and A_t is an increasing process.

Square Integrable Martingales

Definition (Square Integrable Martingale)

A continuous martingale $X = \{X_t, \mathcal{F}_t, 0 \le t < \infty\}$ is said to be square integrable if $\mathbb{E}(X_t)^2 < \infty, \forall t \ge 0$. The class of stochastic processes with this property is denoted by \mathcal{M}_2^c , where the *c* is for continuous.

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Definition (Quadratic Variation)

 $\langle X \rangle_t$ is called the quadratic variation of X_t , where $\langle X \rangle_t$ is the adapted increasing process in the Doob -Meyer Decomposition of X_t^2 . Or also, it is the unique process such that $X_t^2 - \langle X \rangle_t$ is a martingale.

If we consider a partition $0 = t_0 < t_1 < ... < t_n = t$ of [0, t]

$$\langle X \rangle_t = \lim_{n \to \infty} \sum_{k=1}^n (X_{t_k} - X_{t_{k-1}})^2.$$

Square Integrable Martingales

Definition (Cross Variation)

If $X, Y \in \mathcal{M}_2^c$, their cross variation is $\langle X, Y \rangle_t = \frac{1}{4} [\langle X + Y \rangle_t - \langle X - Y \rangle_t].$

And again, it can be thought of as

$$\langle X, Y \rangle_t = \lim_{n \to \infty} \sum_{k=1}^n (X_{t_k} - X_{t_{k-1}}) (Y_{t_k} - Y_{t_{k-1}})$$

Local Martingales

Definition (Stopping and Optional Times)

Let T be a random time. Then T is a stopping time with respect to the filtration $\{\mathcal{F}_t, t \in \Gamma\}$ if the event $\{\omega \in \Omega : T(\omega) \le t\} \in \mathcal{F}_t$ for every $t \ge 0$. A random time is an optional time if $\{\omega \in \Omega : T(\omega) < t\} \in \mathcal{F}_t$ for every $t \ge 0$.

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Definition (Local Martingale)

Let $X = \{X_t, \mathcal{F}_t, 0 \le t < \infty\}$ be a continuous process with $X_0 = 0$ \mathbb{P} -a.s. If there exists a nondecreasing sequence $\{T_n\}_{n=1}^{\infty}$ of stopping times of \mathcal{F}_t such that $\{X_t^{(n)} = X_{t \land T_n}\}$ is a martingale for each $n \ge 1$ and $\mathbb{P}[\lim_{n \to \infty} T_n = \infty] = 1$, then we say that X is a continuous local martingale. It is usually written $X \in \mathcal{M}^{c,loc}$.

Consider the half-line $[0,\infty)$ and divide it in tiny intervals of length δ as it is shown below.



Figure 1: Dividing the half-line

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Each one of these intervals corresponds to a time slot of length δ . Let us assume that we toss a fair coin and create the random variable for each interval [i, i + 1]

$$X_i^{\delta} = egin{cases} \sqrt{\delta}, & ext{with probability 1/2} \ -\sqrt{\delta}, & ext{with probability 1/2} \end{cases}$$

Note that X_i 's are independent and

$$\mathbb{E}[X_i^{\delta}] = 0,$$
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Therefore, we define the process $B_n^{\delta}(t)$ as

$$B_n^{\delta}(t) = B(n\delta) = \sum_{i=1}^n X_i^{\delta}.$$

Then $\forall t \in [0,\infty)$, as $n \to \infty$ and $\delta \to 0$

$$B_n^{\delta} \xrightarrow{dist} B(t) \sim N(0,t)$$

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$$B_n^{\delta} \xrightarrow{dist} B(t) \sim N(0,t)$$

Since the coin tosses are independent, we conclude that B(t) has independent increments, which means that for $0 \le t_1 < t_2 < ... < t_n$, the random variables

$$B(t_2) - B(t_1), B(t_3) - B(t_2), ..., B(t_m) - B(t_{m-1})$$

are independent.

Definition (Brownian Motion)

A standard, one dimensional Brownian Motion is a continuous process $B = \{B_t\}$ adapted to a filtration $\{\mathcal{F}_t, t \in \Gamma\}$ that satisfies the properties

- ▶ $B_0 = 0$ \mathbb{P} -a.s.,
- $B_t B_s$ is independent of \mathcal{F}_s for $0 \le s \le t$,

• $B_t - B_s$ is normally distributed with mean 0 and variance t - s, which means it has independent and stationary increments.

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$$Y_t = \begin{cases} tB_{\frac{1}{t}}, & 0 < t < \infty, \\ 0, & t = 0. \end{cases}$$

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4. (Symmetry)

$$-B = \{-B_t\}.$$

Stochastic Calculus

How can we define

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for a general stochastic process X?

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Definition (Simple Process)

A process S is called simple if there exists a strictly increasing sequence of real numbers $\{t_n\}_{n=0}^{\infty}$ such that $t_0 = 0$ and $\lim_{n\to\infty} t_n = \infty$, as well as a sequence of random variables $\{\epsilon_n\}_{n=0}^{\infty}$ with $\sup_{n\geq 0} |\epsilon_n(\omega)| \leq C$, $\forall \omega \in \Omega$, $\epsilon_n \in \mathcal{F}_{t_n}$, such that

$$S_t(\omega) = \epsilon_0(\omega) \mathbb{1}_{\{0\}}(t) + \sum_{i=0}^{\infty} \epsilon_i(\omega) \mathbb{1}_{\{t_i, t_{i+1}\}}(t).$$

The natural path to proceed is to apply the martingale transform and define the integral of a simple process by

$$I_t(S) = \sum_{i=0}^{\infty} \epsilon_i (M_{t \wedge t_{i+1}} - M_{t \wedge t_i}), \ \ 0 \leq t < \infty.$$

 $I_0(X) = 0$ P-a.s.

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(Linearity) $I(\alpha X + \beta Y) = \alpha I(X) + \beta I(Y) \quad \forall \alpha, \beta \in \mathbb{R}$

(Martingale Property) $\mathbb{E}[I_t(X)|\mathcal{F}_s] = I_s(X) \mathbb{P}$ -a.s.

Itô's Rule

Itô's Rule

Definition (Semimartingale)

A continuous semimartingale $X = \{X_t, \mathcal{F}_t, 0 \le t < \infty\}$ is an adapted process which has the decomposition \mathbb{P} -a.s.

$$X_t = X_0 + M_t + B_t; \quad 0 \le t < \infty$$

where $M = \{M_t, \mathcal{F}_t, 0 \le t < \infty\} \in \mathcal{M}^{c, loc}$ and $B = \{B_t\}$ is the difference of continuous non decreasing, adapted processes

$$B_t = A_t^+ - A_t^-.$$

This decomposition should be the minimal decomposition of B, i.e, A_t^+ and A_t^- are the positive and negative variations of B on [0, t].

Itô's Rule

Theorem (Itô's Rule)

Let $f : \mathbb{R} \to \mathbb{R}$ be a function of class C^2 and let $X = \{X_t, \mathcal{F}_t; 0 \le t < \infty\}$ be a continuous semimartingale. Then \mathbb{P} -a.s.

$$f(X_t) = f(X_0) + \int_0^t f'(X_s) dM_s + \int_0^t f'(X_s) dB_s + \frac{1}{2} \int_0^t f''(X_s) d\langle M \rangle_s,$$

with $0 \le t < \infty$. The above equality can be also be written in differential form:

$$d(f(X_t)) = f'(X_t)dM_t + f'(X_t)dB_t + \frac{1}{2}f''(X_t)d\langle M\rangle_t.$$

Examples

Let $\{B_t : t \ge 0\}$ a Brownian Motion. Then when considering $f(x) = x^2/2$ and $g(x) = x^3/3$ we have from Itô's Rule:

$$\int_0^t B_s dB_s = \frac{1}{2}(B_t^2 - t)$$

and

$$\int_0^t B_s^2 dB_s = \frac{1}{3} B_t^3 - \int_0^t B_s ds.$$

For f, applying the formula above we have

$$\frac{B_t^2}{2} = \frac{B_0^2}{2} + \int_0^t B_s dB_s + \int_0^t d\langle B \rangle_s.$$

Using the fact $B_0 = 0$ Pa.s. and the properties of Brownian Motion we have the result.

Multidimensional Formula

Theorem (Multidimensional Formula) Let $\{M_t = (M_t^{(1)}, ..., M_t^{(d)}\}$ with $M_t^{(i)} \in \mathcal{M}^{c,loc}$ for $1 \le i \le d$ and $\{B_t = (B_t^{(1)}, ..., B_t^{(d)})\}$ a vector of adapted processes of bounded variation with $B_0 = 0\mathbb{P}a.s.$ Let $X_t = X_0 + M_t + B_t$, where $X_0 \in \mathcal{F}_0$. Let $f(t, x) : [0, \infty) \times \mathbb{R}^d \to \mathbb{R}$ be of class $C^{1,2}$. Then, \mathbb{P} -a.s.

$$\begin{split} f(t,X_t) &= f(0,X_0) + \int_0^t \frac{\partial}{\partial t} f(s,X_s) ds + \sum_{i=1}^d \int_0^t \frac{\partial}{\partial x_i} f(s,X_s) dB_s^{(i)} \\ &+ \sum_{j=1}^d \int_0^t \frac{\partial}{\partial x_l} f(s,X_s) dM_s^{(j)} \\ &+ \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d \int_0^t \frac{\partial^2}{\partial x_i \partial x_j} f(s,X_s) d\langle M^{(i)}, M^{(j)} \rangle_s, \quad 0 \le t < \infty. \end{split}$$

Theorem (Levy's Theorem)

Let $X = \{X_t = (X_t^{(1)}, ..., X_t^{(d)}), \mathcal{F}_t, 0 \le t < \infty\}$ be a continuous, adapted process in \mathbb{R}^d such that, for every component $1 \le k \le d$, the process

$$M_t^{(k)} = X_t^{(k)} - X_0^{(k)}, \quad 0 \le t < \infty$$

is a continuous local martingale relative to the filtration $\{\mathcal{F}_t, t \in \Gamma\}$, and whose cross variations are given by

$$\langle M^{(k)}, M^{(j)} \rangle_t = \delta_{kj}t; \quad 1 \le k, j \le d$$

Then $\{X_t, \mathcal{F}_t, 0 \leq t < \infty\}$ is a d-dimensional Brownian Motion.

What we need to prove is that for 0 ≤ s < t, the random vector X_t − X_s is independent of F_s and has a d-variate normal distribution with mean zero and covariance matrix equal to (t − s) × I_d.

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▶ But now, from Levy's continuity theorem, it suffices to prove that for $u \in \mathbb{R}^d$

$$\mathbb{E}[e^{i(u.(X_t-X_s))}|\mathcal{F}_s] = e^{-\frac{1}{2}\|u\|^2(t-s)}, \quad \mathbb{P}-a.s.$$

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$$\mathbb{E}[e^{i(u.(X_t-X_s))}|\mathcal{F}_s] = e^{-\frac{1}{2}||u||^2(t-s)}, \quad \mathbb{P}-a.s.$$

▶ When we fix $u = (u_1, ..., u_d) \in \mathbb{R}^d$, the function $f(x) = e^{i(u.x)}$ satisfies

$$\frac{\partial}{\partial x_j}f(x) = iu_jf(x), \quad \frac{\partial^2}{\partial x_j\partial x_k}f(x) = -u_ju_kf(x)$$

► Therefore, using the multidimensional formula for the real and imaginary parts of f, we obtain $T = \mathcal{D}(f(\mu X)) = \mathcal{D}(f(\mu X)) = 1 \sum_{i=1}^{d} f(\mu X_i) + 1 \sum_{i=1}^{d} f(\mu X_i) + 1 \sum_{i=1}^{d} f(\mu X_i) = 1 \sum_{i=1}^{d} f(\mu X_i) + 1 \sum_{i=1}^{d$

1.
$$\mathcal{R}(e^{i(u.X_t)}) = \mathcal{R}(e^{i(u.X_s)}) - \frac{1}{2} \sum_{j=1}^{d} u_j^2 \int_s^t e^{i(u.X_v)} dv$$

2. $\mathcal{I}(e^{i(u.X_t)}) = \mathcal{I}e^{i(u.X_s)} + \sum_{j=1}^{d} u_j \int_s^t e^{i(u.X_v)} dM_v^{(j)}$

and we have:

$$e^{i(u,X_t)} = e^{i(u,X_s)} + i \sum_{j=1}^d u_j \int_s^t e^{i(u,X_v)} dM_v^{(j)} - \frac{1}{2} \sum_{j=1}^d u_j^2 \int_s^t e^{i(u,X_v)} dv.$$

▶ Note that $|f(x)| \leq 1 \ \forall x \in \mathbb{R}^d$ and since $\langle M^{(j)} \rangle_t = t$, this implies that $M^{(j)} \in \mathcal{M}_2^c$. Hence, the real and imaginary parts of $\{\int_0^t e^{i(u.X_v)} dM_v^{(j)}, \mathcal{F}_t; 0 \leq t < \infty\}$ are not only in $\mathcal{M}^{c,loc}$ but also in \mathcal{M}_2^c .

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- Consequently

$$\mathbb{E}\Big[\int_{s}^{t}e^{i(u,X_{v})}dM_{v}^{(j)}\Big|\mathcal{F}_{s}\Big]=0,\quad \mathbb{P}-a.s.$$

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- Consequently

$$\mathbb{E}\Big[\int_{s}^{t}e^{i(u,X_{v})}dM_{v}^{(j)}\Big|\mathcal{F}_{s}\Big]=0,\quad \mathbb{P}-a.s.$$

▶ Now, for each $A \in \mathcal{F}_s$ and multiplying the last equation by $e^{-i(u,X_s)}\mathbb{1}_A$ and applying expected values we obtain

$$\mathbb{E}[e^{(u.(X_t-X_s))}\mathbb{1}_{A}] = \mathbb{P}(A) - \frac{1}{2} \|u\|^2 \int_{s}^{t} \mathbb{E}[e^{(u.(X_v-X_s))}\mathbb{1}_{A}] dv$$

which is an integral equation already solved with solution :

$$\mathbb{E}[e^{(u.(X_t-X_s))}\mathbb{1}_A] = \mathbb{P}(A)e^{-\frac{1}{2}\|u\|^2(t-s)}, \quad \forall A \in \mathcal{F}_s.$$

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